Parametric g-Formula

Cornell STSCI / INFO / ILRST 3900 Fall 2023 causal3900.github.io

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Learning goals for today

At the end of class, you will be able to

estimate average causal effects with a parametric model

- for the outcome $E(Y | A, \vec{L})$
- for the treatment $P(A \mid \vec{L})$

After class:

▶ Hernán and Robins 2020 Chapter 12.1–12.5, 13, 15.1

Nonparametric estimation

Causal assumptions



Nonparametric estimator

$$\widehat{\mathsf{E}}(Y^{a}) = \frac{1}{n} \sum_{i=1}^{n} \widehat{\mathsf{E}}(Y \mid \vec{L} = \vec{\ell}_{i}, A = a)$$

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For every unit *i*,

- find units who look like them on confounders \vec{L}
- who actually got treatment A = a
- take the average among those units

Then average over all units

Nonparametric estimation breaks down



Nonparametric estimation breaks down

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Parametric estimation: Outcome model

Causal assumptions

$$\vec{L} \xrightarrow{A \to Y} Y$$

Parametric estimator

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First, learn a model to predict Y given \vec{L} and A

$$\hat{\mathsf{E}}(Y \mid \vec{L}, A) = \hat{\alpha} + \vec{L}' \hat{\vec{\gamma}} + A \hat{\beta}$$

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$$\hat{\mathsf{E}}(Y \mid \vec{L}, A) = \hat{\alpha} + \vec{L}' \hat{\vec{\gamma}} + A \hat{\beta}$$

For every unit *i*,

- change the treatment value to a
- predict the outcome

Then average over all units

$$\hat{\mathsf{E}}(Y^{1}) - \hat{\mathsf{E}}(Y^{0}) = \left(\frac{1}{n}\sum_{i=1}^{n} \left(\hat{\alpha} + \hat{\gamma}\ell_{i} + \hat{\beta} \times 1\right)\right) - \left(\frac{1}{n}\sum_{i=1}^{n} \left(\hat{\alpha} + \hat{\gamma}\ell_{i} + \hat{\beta} \times 0\right)\right)$$

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Estimator for the effect $E(Y^1) - E(Y^0)$:

$$\hat{\mathsf{E}}(\mathsf{Y}^{1}) - \hat{\mathsf{E}}(\mathsf{Y}^{0}) = \left(\frac{1}{n}\sum_{i=1}^{n} \left(\hat{\alpha} + \hat{\gamma}\ell_{i} + \hat{\beta} \times 1\right)\right)$$
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With OLS, the parametric g-formula collapses on the coefficient.

The parametric g-formula is more general

Suppose the effect of A depends on L

$$\mathsf{E}(Y \mid A, L) = \alpha + \gamma L + \beta A + \eta A L$$

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$$\hat{\mathsf{E}}(Y^{1}) - \hat{\mathsf{E}}(Y^{0}) = \left(\frac{1}{n}\sum_{i=1}^{n} \left(\hat{\alpha} + \hat{\gamma}\ell_{i} + \hat{\beta} \times 1 + \hat{\eta} \times 1 \times \ell_{i}\right)\right)$$
$$- \left(\frac{1}{n}\sum_{i=1}^{n} \left(\hat{\alpha} + \hat{\gamma}\ell_{i} + \hat{\beta} \times 0 + \hat{\eta} \times 0 \times \ell_{i}\right)\right)$$

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Suppose the effect of A depends on L

$$\mathsf{E}(Y \mid A, L) = \alpha + \gamma L + \beta A + \eta A L$$

Estimator for the effect $E(Y^1) - E(Y^0)$:

$$\begin{split} \hat{\mathsf{E}}(Y^{1}) - \hat{\mathsf{E}}(Y^{0}) &= \left(\frac{1}{n} \sum_{i=1}^{n} \left(\hat{\alpha} + \hat{\gamma}\ell_{i} + \hat{\beta} \times 1 + \hat{\eta} \times 1 \times \ell_{i}\right)\right) \\ &- \left(\frac{1}{n} \sum_{i=1}^{n} \left(\hat{\alpha} + \hat{\gamma}\ell_{i} + \hat{\beta} \times 0 + \hat{\eta} \times 0 \times \ell_{i}\right)\right) \\ &= \frac{1}{n} \sum_{i=1}^{n} \left(\hat{\beta} + \hat{\eta}\ell_{i}\right) \end{split}$$

The g-formula no longer collapses to a coefficient!

Parametric g-formula: Outcome model recap



- 1. Model the outcome mean $E(Y \mid A, \vec{L})$
- 2. Change everyone's treatment to the value of interest
- 3. Predict for everyone
- 4. Take the average

$$\widehat{\mathsf{E}}(Y^{a}) = \frac{1}{n} \sum_{i=1}^{n} \widehat{\mathsf{E}}(Y \mid \vec{L} = \vec{\ell}_{i}, A = a)$$





 $\vec{L} \xrightarrow{A \to Y}$





Propensity score:
$$\pi_i = P(A = A_i | L = L_i)$$



Propensity score: $\pi_i = \mathsf{P}(A = A_i \mid L = L_i)$ Inverse probability weight: $w_i = \frac{1}{\pi_i}$



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$$\pi_i = \mathsf{P}(A = A_i \mid L = L_i)$$
$$w_i = \frac{1}{\pi_i}$$



 L_1

Model the treatment assignment

$$\hat{\mathsf{P}}(\mathsf{A}=1\midec{\mathcal{L}})=\mathsf{logit}^{-1}\left(\hat{lpha}+\hat{ec{\gamma}}ec{\mathcal{L}}
ight)$$

Predict the propensity score for each unit

$$\hat{\pi}_{i} = \begin{cases} \operatorname{logit}^{-1} \left(\hat{\alpha} + \hat{\vec{\gamma}} \vec{L} \right) & \text{if } A_{i} = 1\\ 1 - \operatorname{logit}^{-1} \left(\hat{\alpha} + \hat{\vec{\gamma}} \vec{L} \right) & \text{if } A_{i} = 0 \end{cases}$$

Estimate by inverse probability weighting

$$\hat{\mathsf{E}}(Y^{a}) = \frac{1}{N} \sum_{i:A_{i}=a} \frac{Y_{i}}{\hat{\pi}_{i}}$$

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Two solutions

- 1. Trim the weights
- 2. Truncate the weights

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Both solutions accept bias in order to reduce variance







Drop units with extreme weights



Drop units with extreme weights



Drop units with extreme weights

Changes target population — Biased for full population





Truncate values of extreme weights



Truncate values of extreme weights



Truncate values of extreme weights

Biased: Ignores some confounding